

Multivariate Volatility, Dependence and Copulas

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$$D_r = \int_{-\infty}^{\infty} (x - M_x)^2 \phi(x) dx$$

$$D_r = \sigma^2 = M_x^2 - (M_x)^2$$

$$p_x(\lambda) = \frac{\lambda^x}{x!} e^{-\lambda}$$

$$p(x_1 \leq x \leq x_2) = \int_{x_1}^{x_2} \phi(x) dx$$

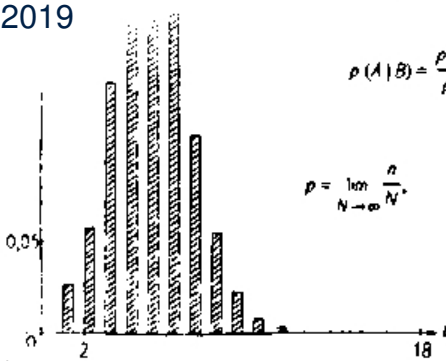
$$M_x = \sum_{i=1}^k p_i x_i$$

$$D_r = \sum_{i=1}^k p_i (x_i - M_x)^2$$

$$\phi(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

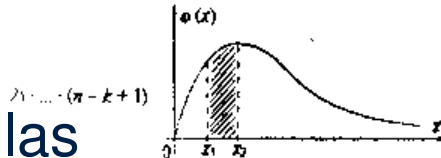
$$\phi(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{x^2}{2\sigma^2}}$$

$$\phi(v) = 4\sqrt{\frac{k^3}{\pi}} v^2 e^{-kv^2}$$



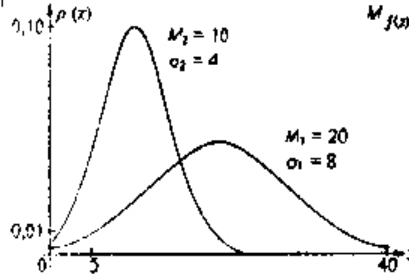
$$p(A|B) = \frac{p(A \cap B)}{p(B)}$$

$$p = \lim_{N \rightarrow \infty} \frac{n}{N}$$



$$\lambda, \dots, (\pi - k + 1)$$

$$x_1, \dots, x_k$$



$$D_r = \int_{-\infty}^{\infty} (x - M_x)^2 \phi(x) dx$$

$$M_x = \int_{-\infty}^{\infty} x \cdot \phi(x) dx$$

$$M_{f(x)} = \int_{-\infty}^{\infty} f(x) \phi(x) dx$$

$$S = v\sigma + \frac{\sigma^2}{2}$$

$$F = G \frac{m_1 m_2}{R^2}$$

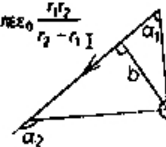
$$f(v) = 4\pi \left(\frac{m_2}{2\pi k^2} \right)^{1/2} v^2 e^{-\frac{mv^2}{2k^2}}$$

$$\phi(\ln x) d(\ln x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} d(\ln x) = \frac{1}{\sqrt{2\pi\sigma x}} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} dx$$

$$\langle v \rangle = \frac{\langle v \rangle t}{n\sqrt{2\pi d^2}}$$

$$C = 4\pi \epsilon_0 \frac{r_1 r_2}{r_2 - r_1}$$

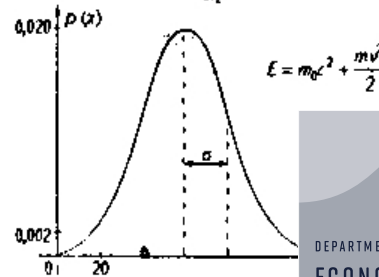
$$C = \frac{\epsilon \epsilon_0 S}{d}$$



$$B = \frac{1}{2ab} (\cos \alpha_1 - \cos \alpha_2)$$

$$A^2 = A_1^2 + A_2^2 + 2A_1 A_2 \cos(\phi_2 - \phi_1)$$

$$hv = A + \frac{mv^2}{2}$$



$$E = mc^2 + \frac{mv^2}{2}$$



- Multivariate Volatility
 - ▶ Simple Models
 - ▶ Dynamic Models
- Realized Covariance
- Dependence
 - ▶ Linear (correlation)
 - ▶ Non-linear
- Copulas

- Portfolio Construction
- Portfolio Sensitivity Analysis
- Value-at-Risk
- Credit Pricing
- Correlation Trading

- Returns \mathbf{r}_t are k by 1
- Use demeaned returns $\boldsymbol{\epsilon} = \mathbf{r}_t - \boldsymbol{\mu}_t$ where $\boldsymbol{\mu}_t$ is conditional mean,
 $\boldsymbol{\mu}_t = \mathbb{E}_{t-1}[\mathbf{r}_t]$
 - ▶ In many cases of interest $\epsilon_t = r_t$
 - ▶ Horizon short and mean small relative to volatility
- Interested in conditional covariance

$$\boldsymbol{\Sigma}_t \equiv \mathbb{E}_{t-1}[\boldsymbol{\epsilon}_t \boldsymbol{\epsilon}_t']$$

- “Devolatilized” residuals

$$u_{i,t} = \epsilon_{i,t} / \sigma_{i,t}, i = 1, 2, \dots, k, \text{ or } \mathbf{u}_t = \boldsymbol{\epsilon}_t \oslash \boldsymbol{\sigma}_t$$

- Standard covariance property

$$\text{Cov} [\mathbf{Az}] = \mathbf{A} \text{Cov} [\mathbf{z}] \mathbf{A}'$$

- Standardized residuals

$$\mathbf{e}_t = \boldsymbol{\Sigma}_t^{-\frac{1}{2}} \boldsymbol{\epsilon}_t$$

- ▶ Devolatilized and decorrelated
- ▶ Since $\text{Cov}_{t-1} [\boldsymbol{\epsilon}_t] = \boldsymbol{\Sigma}_t$, $\text{Cov}_{t-1} \left[\boldsymbol{\Sigma}_t^{-\frac{1}{2}} \boldsymbol{\epsilon}_t \right] = \boldsymbol{\Sigma}_t^{-\frac{1}{2}} \boldsymbol{\Sigma}_t \boldsymbol{\Sigma}_t^{-\frac{1}{2}} = \mathbf{I}_k$

- Conditional Correlation

$$\mathbf{R}_t \equiv \mathbb{E}_{t-1} [\mathbf{u}_t \mathbf{u}_t']$$

- Simplest estimator
- Only reasonable if k is small
- Usually requires $n > k$

Definition (n -period Moving Average Covariance)

The n -period moving average covariance is defined

$$\Sigma_t = n^{-1} \sum_{i=1}^n \epsilon_{t-i} \epsilon'_{t-i}$$

- Use CAP-M or APT to motivate covariance model
- Attribute all covariance to common factors

Definition (n -period Factor Covariance)

The n -period factor covariance is defined as

$$\Sigma_t = \beta' \Sigma_t^f \beta + \Omega_t$$

where $\Sigma_t^f = n^{-1} \sum_{i=1}^n \mathbf{f}_{t-i} \mathbf{f}'_{t-i}$ is the n -period moving covariance of the factors,

$$\beta_t = \left(\sum_{i=1}^n \mathbf{f}_{t-i} \mathbf{f}'_{t-i} \right)^{-1} \sum_{i=1}^n \mathbf{f}_{t-i} \epsilon'_{t-i}$$

is the p by k matrix of factor loadings and Ω_t is a diagonal matrix with $\omega_{j,t}^2 = n^{-1} \sum_{i=1}^n \eta_{j,t-i}^2$ in the j^{th} diagonal position where $\eta_{i,t} = \epsilon_{i,t} - \mathbf{f}'_t \beta_i$ are the regression residuals.

- Is positive semi-definite when $k > p$
- Suitable for large portfolios
- Can be customized in portfolios with different asset classes
- In an Equity – Bond portfolio
 - ▶ 1 factor common to all assets (S&P 500)
 - ▶ 2 factors for equities (Size and Value)
 - ▶ 2 factors for bonds (Curvature and Default Premium)

- Split returns in the orthogonal (uncorrelated) components

$$\min_{\beta, \mathbf{F}} (kT)^{-1} \sum_{i=1}^k \sum_{t=1}^T (y_{i,t} - \mathbf{f}_t \beta_i)^2 \quad \text{subject to } \beta' \beta = \mathbf{I}_k$$

- Solution depends on eigenvalues and eigenvectors, easy to calculate
- Can order factors so that the partial R^2 are decreasing
- Factor 1 explains more than factor 2 which explains more than factor 3, etc.
- Can estimate the number of factors which are common across all assets
- *More details in notes ...*

Definition (n -period Principal Component Covariance)

The n -period principal component covariance is defined as

$$\Sigma_t = \beta_t' \Sigma_t^f \beta_t + \Omega_t$$

where $\Sigma_t^f = n^{-1} \sum_{i=1}^n \mathbf{f}_{t-i} \mathbf{f}'_{t-i}$ is the n -period moving covariance of first p principal component factors, $\hat{\beta}_t$ is the p by k matrix of principal component loadings corresponding to the first p factors, and Ω_t is a diagonal matrix with $\omega_{j,t+1}^2 = n^{-1} \sum_{i=1}^n \eta_{j,t-1}^2$ on the j^{th} diagonal where $\eta_{i,t} = r_{i,t} - \mathbf{f}'_t \beta_{i,t}$ are the residuals from a p -factor principal component analysis.

- Use principal components in place of observable factors
- Same advantage as observable factor covariance
 - ▶ Additional advantage that do not need factors
- Disadvantage that not as easy to implement in a structured setting
- Identical to moving average covariance if all factors used

- Assume that all correlations are identical

$$\sigma_{ij,t} = \rho \sigma_{i,t} \sigma_{j,t}$$

Definition (*n*-period Moving Average Equicorrelation Covariance)

The *n*-period moving average equicorrelation covariance is defined as

$$\Sigma_t = \begin{bmatrix} \sigma_{1,t}^2 & \rho_t \sigma_{1,t} \sigma_{2,t} & \rho_t \sigma_{1,t} \sigma_{3,t} & \dots & \rho_t \sigma_{1,t} \sigma_{k,t} \\ \rho_t \sigma_{1,t} \sigma_{2,t} & \sigma_{2,t}^2 & \rho_t \sigma_{2,t} \sigma_{3,t} & \dots & \rho_t \sigma_{2,t} \sigma_{k,t} \\ \vdots & \vdots & \vdots & \vdots & \vdots \\ \rho_t \sigma_{1,t} \sigma_{k,t} & \rho_t \sigma_{2,t} \sigma_{k,t} & \rho_t \sigma_{3,t} \sigma_{k,t} & \dots & \sigma_{k,t}^2 \end{bmatrix}$$

where $\sigma_{j,t}^2 = n^{-1} \sum_{i=1}^n \epsilon_{j,t}^2$ and ρ_t is estimated using one of the estimators below.

- Moment or maximum likelihood estimator for ρ
- Moment:

$$\begin{aligned} E[\epsilon_{p,t}^2] &= k^{-2} \sum_{j=1}^k \sigma_{j,t}^2 + 2k^{-2} \sum_{o=1}^k \sum_{q=o+1}^k \rho \sigma_{o,t} \sigma_{q,t} \\ &= k^{-2} \sum_{j=1}^k \sigma_{j,t}^2 + 2\rho k^{-2} \sum_{o=1}^k \sum_{q=o+1}^k \sigma_{o,t} \sigma_{q,t} \end{aligned}$$

- Estimator exploits this structure

$$\rho_t = \frac{\sigma_{p,t}^2 - k^{-2} \sum_{j=1}^k \sigma_{j,t}^2}{2k^{-2} \sum_{o=1}^k \sum_{q=o+1}^k \sigma_{o,t} \sigma_{q,t}}.$$

- $\sigma_{j,t}^2$ are the volatilities of the individual assets
- Only appropriate for homogeneous portfolios

- Daily data on S&P 500 constituents from January 1, 1999 – December 31, 2008
- Return only included if present in relevant sample
 - ▶ Full sample
 - ▶ Rolling 252-day sample, centered at sample mid-point
- Full Sample PCA

$k = 194$	1	2	3	4	5	6
Partial R^2	0.263	0.039	0.031	0.023	0.019	0.016
Cumulative R^2	0.263	0.302	0.333	0.356	0.375	0.391

- Full Sample Correlation

Equicorrelation	1-Factor R^2 (S&P 500)	3-Factor R^2 (Fama-French)
0.255	0.236	0.267

Rolling Window Correlations

